

Part 4: Advanced

the really cool stuff...



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In this part I discuss the following advanced topics: Multi-level regressions and SEM Interaction effects in SEM Cluster analysis





Multi-level models

in regression analysis and SEM



Repeated measurements

e.g. participants make 30 decisions

(Partially) within-subjects design

e.g. participants are randomly assigned to 1 of 3 games, and test it once with sound on and once with sound off

Grouped data

e.g. participants perform tasks in groups of 5

A combination of the above





- Consequence: errors are correlated
 - There will be a user-bias (and maybe an task-bias)
- Golden rule: data-points should be **independent**





OK solution...

Take the average of the repeated measurements

- Reduces the number of observations
- It becomes impossible to make inferences about individual tasks/users/etc.







In regression:

- define a random intercept for each user (GLMM)
- impose an error
 covariance structure
 (GEE)







Figs here





Data: 396 participants each make 31 disclosure decisions (binary)

Manipulations:

- Between subjects: 5 justification types: 1:none, 2:useful-foryou, 3:%others, 4:useful-for-others, 5:explanation
- Between subjects: request order (counter-balanced)
- Within subjects: questionID (#1-#31)
- Within subjects: percentage (only for justification types 2, 3 and 4)





What is the effect of the justification types, and does the percentage displayed in the justification play any role?





Naive specification in R, using GLM:

model1 <- glm(decision ~ fmessage*percentage, family=binomial, data=fat2)</pre>

Output:

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)	
(Intercept)	1.585620	0.049354	32.128	< 2e-16	***
fmessage1	-0.218224	0.067850	-3.216	0.00130	**
fmessage2	-0.514137	0.063370	-8.113	4.93e-16	***
fmessage3	-0.630636	0.063346	-9.955	< 2e-16	***
fmessage4	-0.206947	0.067171	-3.081	0.00206	**
percentage	-0.002052	0.001698	-1.209	0.22670	
<pre>fmessage1:percentage</pre>	0.003472	0.002313	1.501	0.13332	
<pre>fmessage2:percentage</pre>	0.006351	0.002176	2.919	0.00351	**
<pre>fmessage3:percentage</pre>	0.003224	0.002175	1.482	0.13830	
<pre>fmessage4:percentage</pre>	0.002145	0.002317	0.926	0.35457	





GLMM = Generalized Linear Mixed-effects Models Works on normal data (LMM) and binary/count data (GLMM)

R package: Ime4

Function: glmer (or lmer)





Random intercept for participant (sessionId):

model2 <- glmer(decision ~ fmessage*percentage + (1|sessionId), family=binomial, data=fat2)





The 15283 data points originate from 493 participant

How do we deal with this?

We could create a separate dummy for each participant-1...

...instead we assume that this intercept is a normally distributed random variable with a certain variance

What are the consequences?

For the between subjects manipulation, standard errors may increase significantly!





Output:

Random effects: Groups Name Variance Std.Dev. sessionId (Intercept) 1.772 1.331 Number of obs: 15283, groups: sessionId, 493

Fixed effects:

	Estimate	Std. Error	z value	Pr(> z)	
(Intercept)	2.023206	0.152286	13.286	< 2e-16	***
fmessage1	-0.200659	0.215050	-0.933	0.350778	
fmessage2	-0.629890	0.204927	-3.074	0.002114	**
fmessage3	-0.708030	0.207900	-3.406	0.000660	***
fmessage4	-0.231913	0.211854	-1.095	0.273657	
percentage	-0.002294	0.001887	-1.216	0.224172	
<pre>fmessage1:percentage</pre>	0.003966	0.002588	1.533	0.125381	
<pre>fmessage2:percentage</pre>	0.008360	0.002422	3.452	0.000556	***
<pre>fmessage3:percentage</pre>	0.003009	0.002453	1.227	0.219986	
<pre>fmessage4:percentage</pre>	0.003125	0.002573	1.215	0.224536	





Can we do better?

Yes; questions are also repeated!

- Again, we could add a dummy variable for each question
- But let's instead add another random intercept

Add random intercept for questionId: model3 <- glmer(decision ~ fmessage*percentage + (1|sessionId) + (1|questionId), family=binomial, data=fat2)





Output:

Random effects: Groups Name Variance Std.Dev. sessionId (Intercept) 4.161 2.040 questionId (Intercept) 2.437 1.561 Number of obs: 15283, groups: sessionId, 493; questionId, 31

Fixed effects:

	Estimate	Std. Error	z value	Pr(> z)	
(Intercept)	2.902810	0.361157	8.038	9.17e-16	***
fmessage1	-0.299082	0.319965	-0.935	0.349926	
fmessage2	-0.951376	0.305720	-3.112	0.001859	**
fmessage3	-1.040422	0.310106	-3.355	0.000793	***
fmessage4	-0.350746	0.315350	-1.112	0.266034	
percentage	-0.001853	0.002304	-0.804	0.421169	
<pre>fmessage1:percentage</pre>	0.003657	0.003150	1.161	0.245569	
<pre>fmessage2:percentage</pre>	0.009889	0.002957	3.344	0.000825	***
<pre>fmessage3:percentage</pre>	0.005157	0.002981	1.730	0.083703	
<pre>fmessage4:percentage</pre>	0.002917	0.003134	0.931	0.351925	





Compare (nested) models with ANOVA:

anova(model2, model3)

Result:

 Df
 AIC
 BIC
 logLik
 deviance
 Chisq
 Chi Df
 Pr(>Chisq)

 model2
 11
 14018
 14102
 -6998.0
 13996
 13996

 model3
 12
 10487
 10578
 -5231.3
 10463
 3533.3
 1
 < 2.2e-16</td>

The difference is significant!





Can we do better?

Maybe percentage has a different influence per participant?

Again, we could add an interaction of percentage*sessionId (lots of dummies!)

But let's instead add a random slope

Add random slope for percentage and sessionId: model3 <- glmer(decision ~ fmessage*percentage + (1+percentage|sessionId) + (1|questionId), family=binomial, data=fat2)





Output:

Random effects: Groups Name Variance Std.Dev. Corr sessionId (Intercept) 4.198e+00 2.048948 percentage 3.344e-05 0.005783 0.20 questionId (Intercept) 2.459e+00 1.568018 Number of obs: 15283, groups: sessionId, 493; questionId, 31

Fixed effects:

	Estimate	Std. Error	z value	Pr(> z)	
(Intercept)	2.915962	0.362914	8.035	9.37e-16	***
fmessage1	-0.300180	0.321299	-0.934	0.350165	
fmessage2	-0.951371	0.307077	-3.098	0.001947	**
fmessage3	-1.040798	0.311430	-3.342	0.000832	***
fmessage4	-0.352260	0.316665	-1.112	0.265963	
percentage	-0.001280	0.002515	-0.509	0.610853	
<pre>fmessage1:percentage</pre>	0.003833	0.003304	1.160	0.246014	
<pre>fmessage2:percentage</pre>	0.010004	0.003113	3.214	0.001311	**
<pre>fmessage3:percentage</pre>	0.005100	0.003132	1.628	0.103431	
<pre>fmessage4:percentage</pre>	0.002877	0.003284	0.876	0.380994	





Compare (nested) models with ANOVA:

anova(model3, model4)

Result:

 Df
 AIC
 BIC
 logLik
 deviance
 Chisq
 Chi Df
 Pr(>Chisq)

 model3
 12
 10487
 10578
 -5231.3
 10463
 10463
 0.3255

 model4
 14
 10488
 10595
 -5230.2
 10460
 2.2451
 2
 0.3255

The difference is **not** significant!





GEE = General Estimating Equations

Works on normal data and binary/count data

R package: geepack

Function: geeglm

Formula:

```
gee <- geeglm(decision ~ fmessage*percentage, id=sessionId,
family=binomial, corstr="exchangeable", data=fat2)
```





The 15283 data points originate from 493 participants, so errors are correlated within each participant

How do we deal with this?

- We allow correlations in the error covariance matrix
- These errors are allowed to correlate with some equal amount *alpha*





Output:

Coefficients:

	Estimate	Std.err	Wald	Pr(> W)	
(Intercept)	1.58530	0.13248	143.19	< 2e-16	***
fmessage1	-0.21784	0.18306	1.42	0.23404	
fmessage2	-0.51499	0.16626	9.59	0.00195	**
fmessage3	-0.63060	0.17529	12.94	0.00032	***
fmessage4	-0.20758	0.17392	1.42	0.23267	
percentage	-0.00174	0.00160	1.18	0.27657	
<pre>fmessage1:percentage</pre>	0.00303	0.00209	2.10	0.14755	
<pre>fmessage2:percentage</pre>	0.00664	0.00217	9.34	0.00224	**
<pre>fmessage3:percentage</pre>	0.00228	0.00203	1.26	0.26228	
<pre>fmessage4:percentage</pre>	0.00239	0.00217	1.21	0.27035	

[...]

Estimated Correlation Parameters: Estimate Std.err alpha 0.202 0.0219 Number of clusters: 493 Maximum cluster size: 31





GLMM can also handle time series data

- Each question is correlated with surrounding questions
- Use "ar1" instead of "exchangeable"
- Specify the order of questions using the "waves" parameter

No examples for this; try it yourself ;-)





Can we do this in SEM too? Yes! Both ways!





Under VARIABLE: Specify id variable (cluster = userid) Under ANALYSIS: Specify complex model (type = complex)





Advantages:

Simple specification, works just like regular SEM

Disadvantages:

Only two levels; no random slopes or double intercepts





Under VARIABLE:

Specify within-subjects variables (within = a b c) Specify between-subjects variables (between = x y z)

Specify id variable (cluster = userid)

Under ANALYSIS:

Specify two-level model (type = twolevel)

Under MODEL:

Specify %within% and %between% effects



GLMM-like SEM

Advantages:

- Can do more than two levels ("threelevel"), and even combine with GEE ("twolevel complex")
- Does intercepts; also random slopes ("twolevel random")
- The random slope can be a dependent variable in another regression (cross-level interactions)

Disadvantages:

- Cannot use categorical indicators
- Can take a long time to estimate (especially "random")





Take a class:

STATS 203

Learn it yourself:

Fitzmaurice, Laird and Ware, "Applied Longitudinal Analysis"

MPlus course videos (the advanced sessions)





Interaction effects in SEM



What is the combined effect of x1 and x2 on y?		x1 = low	x1 = high
Possibilities: Additive effect Super-additive effect	x2 = low	0	5
Sub-additive effect Cross-over	x2 = high	5	10





What is the combined effect of x1 and x2 on y?		x1 = low	x1 = high
Possibilities: Additive effect Super-additive effect	x2 = low	0	5
Sub-additive effect Cross-over	x2 = high	5	15





What is the combined effect of x1 and x2 on y?		x1 = low	x1 = high
Possibilities:			
Additive effect	x2 = low	0	5
Super-additive effect			
Sub-additive effect	x2 = high	5	5
Cross-over			





What is the combined effect of x1 and x2 on y?		x1 = low	x1 = high
Possibilities: Additive effect	x2 = low	0	5
Super-additive effect			
Sub-additive effect Cross-over	x2 = high	5	0





This is easy in regressions Just multiply the dependent variables! y ~ x1*x2

More difficult in SEM

Depends on type of variables: manipulation * manipulation manipulation * factor factor * factor





manipulation * manipulation is easy:

Just create the dummies!

See SEM slides for an example

manipulation * factor:

Multiple groups model or predicted random slopes model

factor * factor:

Predicted random slopes model





"Predicted random slopes model" Pro: Works for all types of variables Con: Cannot use categorical indicators Con: Can take a long time to estimate

- "Multiple groups model"
 - Pro: Easier to estimate
 - Pro: Can sometimes use categorical indicators*
 - Con: Does not work for factor * factor interactions





Under ANALYSIS:

- Specify random slopes (type = random)
- Specify integration (algorithm = integration)

Under MODEL:

- Specify the moderated effect as random: s | y on x;
- Regress the slope on the moderator: s on m;
- Add main effect of moderator: y on m;





Example: is the effect of perceived control on perceived recommendation quality dependent on understandability?

In regression terms:

quality ~ control*underst

In SEM:

s | quality ON control; s ON underst; quality ON underst;





```
ANALYSIS:
  type = random;
  algorithm = integration;
```

```
MODEL:
satisf BY s1* s2-s7;
quality BY q1* q2-q6;
control BY c1* c2-c4;
underst BY u2* u4-u5;
satisf-underst@1;
```

```
satisf ON quality control;
s | quality ON control;
s ON underst;
quality ON underst;
underst ON citem cfriend cgraph;
```











Example: is the effect of perceived control on perceived recommendation quality dependent on the control condition?

In SEM:

s | quality ON control; s ON citem cfriend; quality ON citem cfriend;





```
ANALYSIS:
  type = random;
  algorithm = integration;
```

```
MODEL:
satisf BY s1* s2-s7;
quality BY q1* q2-q6;
control BY c1* c2-c4;
```

```
underst BY u2* u4-u5;
satisf-underst@1;
```

```
satisf ON quality control;
s | quality ON control;
s ON citem cfriend;
quality ON citem cfriend;
underst ON citem cfriend cgraph;
```











Under VARIABLE:

Specify the moderating manipulation as a "grouping" variable: grouping = cctrl(0=none 1=item 2=friend)

- Add a MODEL section for all groups except the baseline Model item:
 - Model friend:

Add corresponding labels to each MODEL to restrict the moderation





```
MODEL:
```

```
satisf BY s1* s2-s7;
quality BY q1* q2-q6;
control BY c1* c2-c4;
underst BY u2* u4-u5;
satisf-underst@1;
```

```
satisf ON quality control (1-2);
quality ON control (p1);
control ON underst (4);
underst ON cgraph (5);
```

```
[satisf] (6);
[quality] (7);
[control] (8);
[underst];
```

```
MODEL item:
satisf ON quality control (1-2);
quality ON control (p2);
control ON underst (4);
underst ON cgraph (5);
```

```
[satisf] (6);
[quality] (7);
[control] (8);
[underst];
```

```
MODEL friend:
  satisf ON quality control (1-2);
  quality ON control (p3);
  control ON underst (4);
  underst ON cgraph (5);
```

```
[satisf] (6);
[quality] (7);
[control] (8);
[underst];
```





Learn it yourself: MPlus course videos (the advanced sessions)





Cluster Analysis

using Latent Categorical Analysis and Mixture Factor Analysis



Putting people into distinct groups...

- ...based on how they answer certain questions
- ...based on behavioral patterns
- ...etc

Two versions:

Based on "raw data": Latent Categorical Analysis Based on factors: Mixture Factor Analysis





Dataset

ID	ltems
1	Wall
2	Status updates
3	Shared links
4	Notes
5	Photos
6	Hometown
7	Location (city)
8	Location (state/province)
9	Residence (street address)
10	Employer
11	Phone number
12	Email address
13	Religious views
14	Interests (favorite movies, etc.)
15	Facebook groups
16	Friend list





Under VARIABLE:

Specify the number of classes: classes = c(2)

Under ANALYSIS:

Specify mixture model: type = mixture

Optionally, specify iterations etc





Under VARIABLE:

Specify the number of classes: classes = c(2)

Under ANALYSIS:

Specify mixture model: type = mixture

Optionally, specify iterations etc (often needed!)

Under MODEL:

Add %overall% and then the factor model

Prepare to wait :-)





Balance the following criteria

- Minimum of BIC
- Maximum entropy
- Loglikelihood levels off
- p-value of successor > .05 (use Lo-Mendell-Rubin adjusted LRT test, available in output: tech4)
- Solution makes sense

Table 9

A comparison of the fit of MFA models with different numbers of classes.

	BIC	Entropy	LL	# of par.	<i>p</i> -Value
1 class 2 classes	16,837 16,578	0.973	- 8277.147 - 8133.179	48 53	0.0069
3 classes	16,442	0.998	-8050.552	58	0.0002
4 classes 5 classes	16,468 16,482	0.998 0.878	- 8048.736 - 8041.459	63 68	0.407 0.999
6 classes 7 classes	16,351 16,359	0.897 0.852	- 7960.902 - 7950.412	73 78	0.812 0.893

The bold values are mentioned in the text as indicators of the optimal number of dimensions.

Fig. 8. Change in loglikelihood between subsequent MFA models.

"It is the mark of a truly intelligent person to be moved by statistics."

George Bernard Shaw